Gaussian Cramer–Rao Bound for Direction Estimation of Noncircular Signals in Unknown Noise Fields

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Abstract—This paper focuses on the stochastic Cramer–Rao bound (CRB) on direction of arrival (DOA) estimation accuracy for noncircular Gaussian sources in the general case of an arbitrary unknown Gaussian noise field parameterized by a vector of unknowns. Explicit closed-form expressions of the stochastic CRB for DOA parameters alone are obtained directly from the Slepian–Bangs formula for general noncircular complex Gaussian distributions. As a special case, the CRB under the nonuniform white noise assumption is derived. Our expressions can be viewed as extensions of the well-known results by Stoica and Nehorai, Ottersten et al., Weiss and Friedlander, Pesavento and Gershman, and Gershman et al. Some properties of these CRBs are proved and finally, these bounds are numerically compared with the conventional CRBs under the circular complex Gaussian distribution for different unknown noise field models.


I. INTRODUCTION

DIRECTION of arrival (DOA) estimation has been an intensive research field since the introduction of so-called high resolution DOA estimation methods. The performance of such methods are often evaluated using the stochastic and deterministic Cramer–Rao bounds (CRBs). Although the deterministic CRB is known not to be achievable in the general case [1], [2], the stochastic CRB can be achieved asymptotically (in the number of measurements) by several high resolution methods, such as stochastic maximum likelihood (ML). The stochastic and deterministic CRBs for the DOA parameters alone have been derived for circular complex Gaussian sources under uniform white noise field in [2]–[5] and [1], respectively. Then, the stochastic CRB has been derived under nonuniform white and arbitrary unknown parametrized noise fields in [6] and [7], respectively. The general case of an arbitrary unknown noise covariance is particularly important in mobile communications because the dominant noise is external in radio frequency systems [8] and consequently its presence introduces correlation between the noise processes of the different sensors and because there is normally no signal-free samples available that could be used for estimating the noise covariance. In these applications, noncircular complex signal with discrete distributions are frequently encountered (e.g., binary phase shift keying and offset quadrature phase shift keying are frequently encountered), but the associated stochastic CRB appears to be prohibitive to compute. Because under rather general conditions, the noncircular complex Gaussian CRB matrix is the largest of all CRB matrices among the class of arbitrary noncircular complex distributions with given covariance matrices (see, e.g., [9, p. 293]), we need an explicit expression of the stochastic CRB under noncircular Gaussian distributions of the sources and arbitrary unknown noise fields which can be used as an upper bound of the stochastic CRB under these discrete distributions. Consequently, this expression appears to be both an extension of results [6] and [7] to general noncircular complex Gaussian distributions and result [10] to nonuniform white and arbitrary unknown parametrized noise fields.

In this paper, we derive closed-form expressions of this stochastic CRB for DOA parameters alone. Our derivation is inspired by the proof presented in [6], [7] applied to the extended Slepian–Bangs formula [10]. But, due to the noncircularity of the sources, the key point of this proof, i.e., that the number of terms of the extended source covariance matrix is equal to the number of real and imaginary parts of both sources covariance matrices, is not valid. Consequently, to retain the main features of the proof given in [6], [7], we must first prove that the stochastic CRB for the DOA parameter is insensitive to the constraints on this extended covariance matrix. This points will be derived from the study of the ML DOA estimation.

This paper is organized as follows. The array signal model is formulated in Section II. Based on this model, explicit closed-form expressions of the stochastic CRBs are derived in Section III. Then we consider parameter identifiability and analyze some properties of the obtained bounds in Section IV. Finally, Section V presents some illustrative examples, and Section VI concludes the paper.

The following notations are used throughout the paper. Matrices and vectors are represented by bold upper case and bold lower case characters, respectively. Vectors are by default in column orientation, while $T$, $H$ and $s$ stand for transpose, conjugate transpose and conjugate respectively. Symbols $\odot$ and $\otimes$ stand for the Hadamard and the Kronecker product, respectively. $\text{vec}(\cdot)$ is the vectorization operator that turns a matrix into a vector by stacking the columns of the matrix one below another, the vec-permutation matrix $K$ transforms $\text{vec}(\cdot)$ to $\text{vec}(\cdot^T)$ for any square matrix and $\bot$ is the ortho-complement of a projector matrix. $\text{Tr}(\cdot)$, $\text{Det}(\cdot)$, $\ln(\cdot)$, $\Re(\cdot)$ and $\Im(\cdot)$ denote the trace, the
determinant, the logarithm, the real and the imaginary part operators, respectively.

II. ARRAY SIGNAL MODEL

Let an array of $M$ sensors receive $K$ ($K < M$) narrowband signals impinging from the sources with unknown DOAs. The $M \times 1$ array snapshot complex vectors can be modeled as

$$y_t = A(\theta)s_t + n_t, \quad t = 1, \ldots, T$$

where $A(\theta) = [a_1, \ldots, a_K]$ is the full column rank steering matrix where each vector $a_k$ is parameterized by the scalar DOA parameter $\theta_k$ and $\theta \triangleq (\theta_1, \ldots, \theta_K)^T$. $s_t = (s_{t,1}, \ldots, s_{t,K})^T$ and $n_t$ model signals transmitted by $K$ sources and additive measurement noise respectively. $s_t$ and $n_t$ are multivariate independent, complex zero-mean. $n_t$ is assumed circular complex Gaussian, spatially uncorrelated or correlated with unknown covariance matrix $E(n_n n_n^H) = Q_n$, while $s_t$ is noncircular complex Gaussian, and possibly spatially correlated or even coherent with $R_s \triangleq E(s_s s_s^H)$ and $R'_{n} \triangleq E(n_s n_s^T)$. This leads to the covariance matrices of $y_t$

$$R_y(\alpha) = AR_sA^H + Q_n \quad \text{and} \quad R'_{y}(\alpha) = AR_s' A^T$$

where the vector $\alpha$ of unknown real parameters collects the DOAs and nuisance parameters. These covariance matrices are classically estimated by $R_y,_{T} \triangleq \frac{1}{T} \sum_{t=1}^{T} y_t y_t^H$ and $R'_{y},_{T} \triangleq \frac{1}{T} \sum_{t=1}^{T} y_t y_t^T$, respectively. Let us consider the following general noise model introduced in [13] and used in [7]:

$$Q_n = Q_n(\sigma)$$

where $\sigma \triangleq (\sigma_1, \ldots, \sigma_K)^T$ is the vector of real unknown coefficients which are used to parameterize the noise covariance matrix. If no a priori information is available concerning the spatial covariance of the sources, $R_s$ and $R'_s$ are generically parameterized by the real parameters

$$\rho \triangleq (\Re([R_s]_{i,j})), (\Im([R_s]_{i,j})), (\Re([R'_s]_{i,j})), (\Im([R'_s]_{i,j})), 1 \leq j < i \leq K$$

$$([R_s]_{i,i}, \Re([R'_s]_{i,i})), (\Im([R'_s]_{i,i})), (\Re([R'_s]_{i,i})), 1 \leq i \leq K).$$

Thus, the vector of $L = K + K^2 + K(K + 1) + N$ unknown real parameters can be written as $\alpha \triangleq (\theta^T, \rho^T, \sigma^T)^T$.

This parameter is supposed identifiable from $(R_y(\alpha), R'_y(\alpha))$, in the following sense:

$$R_y(\alpha) = R_y(\alpha') \quad \text{and} \quad R'_y(\alpha) = R'_y(\alpha') \Rightarrow \alpha = \alpha'.$$

The probability density function (pdf) of $y_t$ presented in the case of uniform white noise in [10] is preserved, and expressed as a function of $\hat{y}_t \triangleq \frac{y_t}{\sqrt{y_t^H y_t}}$ as

$$p(\hat{y}_t) = (\pi)^{-M} [\Det(R_\theta(\alpha))]^{-1/2} \exp \left\{ -\frac{1}{2} \hat{y}_t^H R_\theta^{-1}(\alpha) \hat{y}_t \right\}$$

with

$$R_\theta(\alpha) \triangleq \begin{bmatrix} R_{s} & R_{s}' \\ R_{s}' & R_{s}'' \end{bmatrix}, \quad A(\theta) \triangleq \begin{bmatrix} A(\theta) & O \\ O & A^*(\theta) \end{bmatrix}$$

and

$$Q_n \triangleq \begin{bmatrix} Q_n & 0 \\ 0 & Q_n' \end{bmatrix}.$$

III. STOCHASTIC CRBs

To derive the stochastic CRB of the parameter $\theta$ alone, two approaches could be considered. One of them consists in computing the asymptotic covariance matrix of the ML estimator, and the other is obtained directly from the extended Slepian–Bangs formula derived in [10]. The first approach has been successfully used in the case of uniform white noise fields in [2] and [10], where a closed-form expression of the log-likelihood function concentrated with respect to the full set of the signal and noise nuisance parameters was available [11]. In the case of nonuniform white and linearly parameterized noise fields, such property has appeared to be impossible to obtain in [6] and [14], respectively. Consequently, we concentrate on the second approach. To adapt the proofs given in the circular Gaussian case in [3], [6] and [7] in the uniform white, nonuniform white and arbitrary unknown parameterized noise field respectively, to the noncircular case, the key point $\text{vec}(R_\theta) = Jp$ where $J$ is a constant nonsingular matrix must be preserved. Because $R_\theta$ is structured [see (2.3)], we must first prove that the stochastic CRB for the DOA parameter is insensitive to the constraints on $R_s$. To prove such a property, we consider the ML estimation of $(\theta, \rho, \sigma)$.

A. Maximum Likelihood Estimation

We first note that the log-likelihood function associated with the PDF (2.2) can be classically written (see e.g., [12] and [2]) after dropping the constants as

$$L(\theta, \rho, \sigma) = -\frac{T}{2} \left( \ln[\Det(R_\theta)] + \Tr(R_\theta^{-1} R_{\theta,\mathcal{T}}) \right)$$

with $R_{\theta,\mathcal{T}} \triangleq \frac{1}{T} \sum_{t=1}^{T} \hat{y}_t \hat{y}_t^H$ where the parameter $(\theta, \rho, \sigma)$ is embedded in the covariance matrix $R_\theta$. Due to the structures of $R_s$ and $Q_n$ in $R_\theta$ (see (2.3)), the ML estimation of $(\theta, \rho, \sigma)$ becomes a constrained optimization problem which is not standard. Despite this difficulty, we prove in Appendix A in the following lemma.

Lemma 1: If the sample covariance matrix $R_{\theta,\mathcal{T}}$ is positive definite, the joint constrained and unconstrained ML estimates which maximize the log-likelihood function (3.1) coincide.

B. Stochastic Cramer–Rao Bound Expressions

From the previous lemma, the stochastic CRB for the signal DOAs associated with the constrained and unconstrained array signal models coincide. Using the unconstrained model, let $\alpha = (\theta^T, \rho^T, \sigma^T)^T$ with here $\rho$ contains the $4K^2$ real parameters of the unconstrained matrices $R_s$, i.e.,

$$\rho \triangleq (\Re([R_s]_{i,j})), (\Im([R_s]_{i,j})), (\Re([R'_s]_{i,j})), (\Im([R'_s]_{i,j})), 1 \leq j < i \leq 2K; (\Re([R'_s]_{i,i})), 1 \leq i \leq 2K).$$

With this unconstrained model, we can follow along the lines of the derivation given in [7] where $R_y = AR_sA^H + Q_n$
is replaced here by \( R_\theta = A R_\theta A^H + Q_\theta \) because the key point of the derivation, i.e., the relation \( \text{vec}(R_\theta) = J \rho \) where \( J \) is a constant nonsingular complex matrix is preserved. By adapting the proof given in [7], the following result is proved in Appendix B.

Result 1: The normalized (i.e., for \( T = 1 \)) DOA-related block of CRB for noncircular complex Gaussian (NCG) sources in the presence of an arbitrary unknown (AU) noise field is given by the following explicit expression:

\[
\text{CRB}_\text{AU}^{\text{NCG}}(\theta) = \frac{1}{2} \left\{ \Re \left[ (D^H \Pi_{\tilde{A}} D) \right] \right. \\
\left. \circ \left( [R_\theta \tilde{A}^H, R_\theta^* A^T] R_\theta^{-1} \left[ \tilde{A} R_\theta \tilde{A}^* R_\theta^* \right]^T \right) \right. \\
\left. - M_\text{AU}^{\text{NCG}} T_\text{AU}^{\text{NCG}}^{-1} M_\text{AU}^{\text{NCG}T} \right\} \quad (3.2)
\]

using real matrices

\[
M_\text{AU}^{\text{NCG}} = 2 \Re \left\{ \left[ D^H \Pi_{\tilde{A}} \right] \circ \left( G \tilde{A} R_\theta \tilde{A}^* R_\theta^* \right)^T \right\} \\
+ 2 \Re \left\{ \left[ D^H \Pi_{\tilde{A}} \right] \circ \left( G' \tilde{A}^* R_\theta^* \right)^T \right\} \\
- M_\text{AU}^{\text{NCG}} T_\text{AU}^{\text{NCG}}^{-1} M_\text{AU}^{\text{NCG}T} \quad (3.3)
\]

\[
T_\text{AU}^{\text{NCG}} = 4 \Re \left\{ P^H \left[ G^T \circ \Pi_{\tilde{A}} \right] P \right\} - 2 \left\{ P^H \left[ (\Pi_{\tilde{A}})^T \circ \Pi_{\tilde{A}} \right] P \right\} \quad (3.4)
\]

with

\[
Q \triangleq \left\{ \text{vec}(e_1 e_2^T), \text{vec}(e_2 e_3^T), \ldots, \text{vec}(e_k e_{k+1}^T) \right\} \\
P \triangleq \left\{ \text{vec}(Q_1), \text{vec}(Q_2), \ldots, \text{vec}(Q_n) \right\}
\]

where the vector \( e_i \) contains one in the \( i \)th position and zeros elsewhere and \( Q_k \triangleq \text{vec}(Q_k) \). \( A \triangleq d(A) / d \theta \). \( D \triangleq [d_1, \ldots, d_K] \). \( \tilde{A} \triangleq Q_1^{-1/2} R_\theta Q_1^{-1/2} A. \tilde{D} \triangleq d(A) / d \theta \). \( \tilde{R}_\theta \triangleq Q_1^{-1/2} R_\theta Q_1^{-1/2} \). \( \tilde{R}_y \triangleq Q_1^{-1/2} R_y Q_1^{-1/2} \). \( \tilde{G} \triangleq (G_y - \tilde{R}_y) (\tilde{R}_y^{-1} \tilde{R}_y^*)^{-1} \) and \( G \triangleq -\tilde{G} \).

When the noise is spatially uncorrelated with different sensor noise variances (nonuniform white noise), i.e., \( Q_k(\sigma) = \text{Diag}(\sigma_1^2, \ldots, \sigma_K^2) \). Result 1 takes the following form that is proved in Appendix C.

Result 2: For noncircular complex Gaussian sources, the normalized DOA-related block of CRB under the nonuniform (NU) white noise assumption is given by the following explicit expression:

\[
\text{CRB}_\text{NU}^{\text{NCG}}(\theta) = \frac{1}{2} \left\{ \Re \left[ (D^H \Pi_{\tilde{A}} D) \right] \right. \\
\left. \circ \left( [R_\theta \tilde{A}^H, R_\theta^* A^T] R_\theta^{-1} \left[ \tilde{A} R_\theta \tilde{A}^* R_\theta^* \right]^T \right) \right. \\
\left. - M_\text{NU}^{\text{NCG}} T_\text{NU}^{\text{NCG}}^{-1} M_\text{NU}^{\text{NCG}T} \right\} \quad (3.5)
\]

using real matrices

\[
M_\text{NU}^{\text{NCG}} = 2 \Re \left\{ \left[ D^H \Pi_{\tilde{A}} \right] \circ \left( G \tilde{A} R_\theta \tilde{A}^* R_\theta^* \right)^T \right\} \\
+ 2 \Re \left\{ \left[ D^H \Pi_{\tilde{A}} \right] \circ \left( G' \tilde{A}^* R_\theta^* \right)^T \right\} \\
- M_\text{NU}^{\text{NCG}} T_\text{NU}^{\text{NCG}}^{-1} M_\text{NU}^{\text{NCG}T} \quad (3.6)
\]

\[
T_\text{NU}^{\text{NCG}} = 2 \left\{ G^T \circ \left( \Pi_{\tilde{A}} \right)^T \circ \left( \Pi_{\tilde{A}} \right) \right\} \quad (3.7)
\]

C. Single Source Case

In the particular case of one signal source, it is shown in Appendix D that the stochastic CRB given by Result 2 can be simplified to

Result 3: The CRB of \( \theta_1 \) for a noncircular complex Gaussian source corrupted by nonuniform white noise field is given by the expression

\[
\text{CRB}_\text{NU}^{\text{NCG}}(\theta_1) = \frac{1}{\sigma_1^2} \left[ \frac{2 \sigma_1^2}{|a_1|^2} - \sigma_1^2 \right] + (1 - \sigma_1^2) \left( \frac{|a_1|^2}{|a_1|^2 r_1^2} \right) \quad (3.8)
\]

where the noncircularity rate \( \rho_1 \) is defined by \( E[\sigma_1^2] = \rho_1 e^{\sigma_1^2} E[\sigma_1^2] \) and satisfies \( 0 \leq \rho_1 \leq 1 \). The SNR is defined as in [6, rel. (48)] by \( r_1 \triangleq \sigma_1^2 / \sigma_1^2 \). The noise dependent factor \( (2|a_1|^2) / |a_1|^2 \). \( \Pi_{\tilde{A}}^{-1} d_1 = 2M\sum_{i=1}^{M} \sigma_i^2 \). The noise vector \( d_1 = d_{\theta_1} / d_{\theta_1} \).

Expression (3.8) is similar to those given in the uniform white noise case [10, rel. (3.14)], except that here, the term \( \sigma_1^2 \) given in Result 3 is not a purely geometric factor. Consequently, similarly as in the uniform white noise case [10, Result 5], the stochastic CRB decreases monotonically as the SNR increases and as the noncircularity rate \( \rho_1 \) increases from \( \text{CRB}_\text{NU}^{\text{NCG}}(\theta_1) = (1/\sigma_1^2) \). \( (1 + 1/|a_1|^2 r_1^2) \) (\( \rho_1 = 0 \). In the case of \( \text{CRB}_\text{NU}^{\text{NCG}}(\theta_1) = (1/\sigma_1^2) \). \( (1 + 1/|a_1|^2 r_1^2) \) (\( \rho_1 = 1 \)).

We note that an expression of \( \text{CRB}_\text{NU}^{\text{NCG}}(\theta_1) \) has already been given [6, rel. (46)], but with a more intricate expression.

We note, that Result 1 cannot be simplified in the case of one signal source because the expression of the product \( M_\text{NU}^{\text{NCG}T} T_\text{NU}^{\text{NCG}}^{-1} M_\text{NU}^{\text{NCG}} \) is very intricate.

IV. DISCUSSION

In this section, we consider parameter identifiability and comparisons of our derived expressions to well known results on the stochastic and deterministic CRBs.

A. Parameter Identifiability

An issue of great importance is the parameter identifiability. That is, how many parameters can be uniquely estimated from \( R_y \) and \( R_y' \). Naturally, the necessary condition (see, e.g., [7, rel. (45)]) derived from the non singularity of the Fisher information matrix associated with the unconstrained model, i.e.,
\[ K + (2K)^2 + N \leq (2M)^2 \text{ is not sufficient. For example, a more severe necessary condition is that the number of unknown real-valued parameters is no more than the number of independent equations. It follows from (2.1), that this is equivalent to} \]

\[ K + K^2 + K(K+1) + N \leq M^2 + M(M+1). \]

Consequently in the case of noncircular signals, the number \( N \) of real-valued unknown parameters is no more than the number of independent equations. It follows from (2.1), that this is equivalent to

\[ K + K^2 + K(K+1) + N \leq M^2 + M(M+1). \]

**B. Comparisons Between CRBs**

Consider the situation when the source signals are circular complex Gaussian (CG) while yet, not using this, therefore \( \mathbf{R}_s' = \mathbf{O} \) and consequently \( \mathbf{R}_y = \begin{bmatrix} \mathbf{R}_y & \mathbf{O} \\ \mathbf{O} & \mathbf{R}_y \end{bmatrix}, \mathbf{G}' = \mathbf{O} \) and \( \mathbf{G} = \mathbf{G}_y^{-1} \). Then our general stochastic CRB formulae given by Results 1 and 2 simplify to the well known bounds in the circular case

\[
\text{CRB}_{\text{CU}}^{\text{CG}}(\theta) = \frac{1}{2} \left\{ \mathbb{R} \left( \begin{bmatrix} \mathbf{D}^H \mathbf{R}_y^{-1} \mathbf{A} \end{bmatrix} \mathbf{D} \right) \right\}^{-1} - \mathbf{M}_{\text{AU}}^{\text{CG}T} \mathbf{M}_{\text{AU}}^{\text{CG}}^{-1} - \mathbf{M}_{\text{AU}}^{\text{CG}T} \mathbf{M}_{\text{AU}}^{\text{CG}}^{-1} \right\}^{-1}
\]

where

\[
\mathbf{M}_{\text{AU}}^{\text{CG}} = 2 \mathbb{R} \left\{ \mathbf{Q}^T \left( \begin{bmatrix} \mathbf{D}^H \mathbf{R}_y^{-1} \mathbf{A} \end{bmatrix} \mathbf{D} \right) \right\}
\]

\[
\mathbf{T}_{\text{AU}}^{\text{CG}} = 4 \mathbb{R} \left\{ \mathbf{P}^H \left( \begin{bmatrix} \mathbf{R}_y^{-1} \mathbf{A} \end{bmatrix} \mathbf{D} \right) \right\}
\]

\[
-2 \left( \mathbf{P}^H \left( \begin{bmatrix} \mathbf{R}_y^{-1} \mathbf{A} \end{bmatrix} \mathbf{D} \right) \right)
\]

(4.1)

in the case of arbitrary unknown noise field and

\[
\text{CRB}_{\text{CU}}^{\text{CG}}(\theta) = \frac{1}{2} \left\{ \mathbb{R} \left( \begin{bmatrix} \mathbf{D}^H \mathbf{R}_y^{-1} \mathbf{A} \end{bmatrix} \mathbf{D} \right) \right\}^{-1} - \mathbf{M}_{\text{NU}}^{\text{CG}T} \mathbf{M}_{\text{NU}}^{\text{CG}}^{-1} - \mathbf{M}_{\text{NU}}^{\text{CG}T} \mathbf{M}_{\text{NU}}^{\text{CG}}^{-1} \right\}^{-1}
\]

where

\[
\mathbf{M}_{\text{NU}}^{\text{CG}} = 2 \mathbb{R} \left\{ \left( \begin{bmatrix} \mathbf{D}^H \mathbf{R}_y^{-1} \mathbf{A} \end{bmatrix} \mathbf{D} \right) \right\}
\]

\[
\mathbf{T}_{\text{NU}}^{\text{CG}} = 2 \left( \mathbf{R}_y^{-1} \mathbf{R}_y^{-1} - \left( \begin{bmatrix} \mathbf{R}_y^{-1} \mathbf{A} \end{bmatrix} \mathbf{D} \right) \right)
\]

(4.2)

in the case of nonuniform white noise field, both derived from the circular complex Slepian–Bangs formula in [7] and [6], respectively. Consequently

\[
\text{CRB}_{\text{NU}}^{\text{CG}}(\theta) |_{\text{CG}} = \text{CRB}_{\text{CU}}^{\text{CG}}(\theta)
\]

and

\[
\text{CRB}_{\text{NU}}^{\text{CG}}(\theta) |_{\text{CG}} = \text{CRB}_{\text{CU}}^{\text{CG}}(\theta).
\]

Next, let us consider the situation when the noise is uniform and spatially white (U) while yet, not using this, the noise is modeled using \( N > 1 \) parameters. Comparing (3.2) and (3.5) under these conditions with

\[
\text{CRB}_{\text{NU}}^{\text{CG}}(\theta) = \sigma_n^2 \left\{ \left( \mathbf{D}^H \mathbf{R}_y^{-1} \mathbf{A} \end{bmatrix} \mathbf{D} \right) \right\}^{-1} - \mathbf{M}_{\text{NU}}^{\text{CG}T} \mathbf{M}_{\text{NU}}^{\text{CG}}^{-1} - \mathbf{M}_{\text{NU}}^{\text{CG}T} \mathbf{M}_{\text{NU}}^{\text{CG}}^{-1} \right\}^{-1}
\]

we obtain

\[
\text{CRB}_{\text{NU}}^{\text{CG}}(\theta) |_{\text{CG}} = \text{CRB}_{\text{CU}}^{\text{CG}}(\theta)
\]

and

\[
\text{CRB}_{\text{NU}}^{\text{CG}}(\theta) |_{\text{CG}} = \text{CRB}_{\text{CU}}^{\text{CG}}(\theta).
\]

Obtained in [10], we have because \( \mathbf{M}_{\text{AU}}^{\text{CG}T} \mathbf{M}_{\text{AU}}^{\text{CG}}^{-1} \mathbf{M}_{\text{AU}}^{\text{CG}T} \) is nonnegative definite

\[
\text{CRB}_{\text{AU}}^{\text{CG}}(\theta) |_{\text{CG}} \geq \text{CRB}_{\text{CU}}^{\text{CG}}(\theta)
\]

and

\[
\text{CRB}_{\text{NU}}^{\text{CG}}(\theta) |_{\text{CG}} \geq \text{CRB}_{\text{CU}}^{\text{CG}}(\theta).
\]

For uniform white noise field and now if this a priori knowledge is taken into account \( \mathbf{Q}_n = \sigma_n^2 \mathbf{I}_M \). Hence \( \mathbf{M}_{\text{AU}}^{\text{CG}} \) becomes an \( K \times 1 \) vector \( \begin{bmatrix} m_1, \ldots, m_K \end{bmatrix}^T \) and \( \mathbf{Q}_n = \mathbf{I}_M \). From (B.5), we obtain

\[
m_k = 2 \mathbb{R} \left\{ \mathbf{T} \left( \begin{bmatrix} \mathbf{D}^H \mathbf{R}_y^{-1} \mathbf{A} \end{bmatrix} \mathbf{D} \right) \right\} = 0
\]

because the column space of \( \mathbf{R}_y^{-1} \mathbf{A} \) is generated by the column space of \( \mathbf{A} \). Therefore, as expected, our general stochastic CRB formulas (3.2) and (3.5) simplify to (4.3) in the uniform white case.

The next result compares the CRBs \( \text{CRB}_{\text{CU}}^{\text{CG}}(\theta) \) [resp. \( \text{CRB}_{\text{CU}}^{\text{CG}}(\theta) \)] and \( \text{CRB}_{\text{CU}}^{\text{CG}}(\theta) \) [resp. \( \text{CRB}_{\text{CU}}^{\text{CG}}(\theta) \)] associated with Gaussian sources with the same first covariance matrix \( \mathbf{R}_s \) and the same arbitrary noise covariance matrix \( \mathbf{Q}_n \)

\[
\text{CRB}_{\text{CU}}^{\text{CG}}(\theta) \leq \text{CRB}_{\text{CU}}^{\text{CG}}(\theta).
\]

**Proof:** For \( \mathbf{Q}_n = \sigma_n^2 \mathbf{I}_M \), this result is proved in [10] and for a single source corrupted by nonuniform white noise, this result is a consequence of Result 3. For an arbitrary number of sources and an arbitrary noise covariance matrix \( \mathbf{Q}_n \), we have to elaborate a little bit because the structure of the matrices \( \mathbf{M} \) and \( \mathbf{T} \) (see (3.3) (3.4), (3.6) (3.7) and (4.1) (4.2)) are very intricate.

Consider the ML estimate of \( \theta \) derived under the circular Gaussian distribution for an arbitrary noise covariance matrix \( \mathbf{Q}_n \). By the asymptotic efficiency of the ML estimator, its asymptotic covariance matrix \( \text{Cov}_{\text{ML}}^{\text{CG}}(\theta) \) satisfies:

\[
\text{Cov}_{\text{ML}}^{\text{CG}}(\theta) = \text{CRB}_{\text{ML}}^{\text{CG}}(\theta).
\]

This ML.CG estimator is given by a second-order algorithm built from \( \mathbf{R}_y^T \). To proceed, we need the following lemma that is proved in Appendix E.

**Lemma 2:** All DOAs consistent estimates given by second-order algorithms based on \( \mathbf{R}_y^T \) only, that do not suppose explicitly the sources to be spatially uncorrelated, are robust to the
noncircularity of the Gaussian sources, i.e., the asymptotic performances are those of the standard complex circular Gaussian case.

Consequently, if the MLCG estimator is used under noncircular Gaussian sources with the same noise covariance $Q_n$, the asymptotic covariance matrix $\text{Cov}_{Q_n}^{\text{ML-MLCG}}(\theta)$ of its DOAs estimate which is based on $R_\theta T$ only, is preserved.

$$\text{Cov}_{Q_n}^{\text{ML-MLCG}}(\theta) = \text{Cov}_{Q_n}^{\text{ML-MLCG}}(\theta).$$

This observation, along with the general CRB inequality, implies that

$$\text{Cov}_{Q_n}^{\text{ML-MLCG}}(\theta) \geq \text{CRB}_{Q_n}^{\text{ML-MLCG}}(\theta)$$

and the proof is complete.

Let us now compare the stochastic and asymptotic deterministic CRBs in the case of colored or nonuniform white noise field for noncircular complex source signals. First, we note that the following expression of the asymptotic deterministic CRB proved in [6] in the circular case remains valid in the noncircular case as well

$$\text{CRB}_{Q_n}^{\text{DET}}(\theta) = \frac{1}{2} \{ \mathbf{R} \left[ (\mathbf{D}^H \mathbf{M} \mathbf{A}^H) \otimes \mathbf{R}_\sigma^T \right] \}^{-1}. \quad (4.5)$$

We prove in Appendix F, the following result.

Result 5: If $\mathbf{R}_8$ is nonsingular

$$\text{CRB}_{Q_n}^{\text{DET}}(\theta) \leq \text{CRB}_{Q_n}^{\text{NC}}(\theta).$$

We note that $\mathbf{R}_8$ may be singular when $\mathbf{R}_6$ and $\mathbf{R}_6'$ are nonsingular, for example when the sources are uncorrelated with noncircularity rate equals to 1. We prove in Appendix E that this result extends to this particular case.

V. ILLUSTRATIVE EXAMPLES

The purpose of this section is to illustrate Results 1-4 and to compare these stochastic CRBs to the stochastic CRBs under circular complex Gaussian distributed source signals as well with the deterministic CRB. We consider throughout this section one or two independent and equipowered sources with identical noncircularity rate ($\rho_1 = \rho_2$). These sources impinge on a uniform linear array of $M = 10$ sensors separated by a half-wavelength for which $\alpha_k = (1, e^{i \theta_k}, \ldots, e^{i(M-1)\theta_k})^T$ where $\theta_k = \pi \sin(\alpha_k)$ with $\alpha_k$, the DOAs relative to the normal of array broadside. We assume that the noise field is modeled by the three following covariance matrices $Q_{\ell_i}^{(i)}$, $i = 1, 2, 3$. The first two models and the third model come from [7] and [6], respectively.

$$Q_{\ell_i}^{(1)}(k, \ell) = \sigma^2 \exp(-|k - \ell|^2 \zeta),$$
$$Q_{\ell_i}^{(2)}(k, \ell) = \sigma^2 \exp(-|k - \ell| \zeta),$$
$$Q_{\ell_i}^{(3)} = \text{Diag}(\sigma^2_1, \ldots, \sigma^2_M).$$

In the first two colored noise field models, $\mathbf{\sigma}^T = (\sigma^2_1, \zeta)$ where $\zeta$ is the “color” parameter and the signal-to-noise (SNR) is defined by $\sigma^2_1 / \sigma^2_n$ and in the nonuniform white noise field model

$$\mathbf{\sigma}^T = (\sigma^2_1, \sigma^2_2, \ldots, \sigma^2_{10})$$
$$= (10, 0.2, 1.5, 0.5, 8.0, 0.7, 1.1, 3.0, 6.0, 3.0)$$

and the SNR is defined by $\sigma^2_n / 10 \sum_{i=1}^{10} 1 / \sigma^2_i$.

In Figs. 1 and 2, we consider two sources and compare the stochastic CRBs under circular and noncircular (with $\rho_1 = \rho_2 = 1$) complex Gaussian distributed source signals to the deterministic CRB. The first two and the third noise field models are used in Figs. 1 and 2 respectively. Fig. 1 shows the bounds $\text{CRB}_{\text{AU}}^{\text{NC}}(\theta_1)$, $\text{CRB}_{\text{AU}}^{\text{NC}}(\theta_1)$ and $\text{CRB}_{\text{AU}}^{\text{DET}}(\theta_1)$ plotted against $\zeta$ for $\Delta \theta = \theta_2 - \theta_1 = 0.1$ rd and $\Delta \phi = 0.52$ rd. Compared to

$^1$All the CRBs are computed for $T = 1$. Note that you find by simulation that the different CRBs depend on $\theta_1, \theta_2, \phi_1, \phi_2$ by only $\Delta \theta = \theta_2 - \theta_1$ and $\Delta \phi = \phi_2 - \phi_1$ for two equipowered sources with identical noncircularity rates.
Fig. 3. Ratio \( r \triangleq \frac{\text{CRB}_{\text{NU}}^{\text{NGC}}(\theta_1)}{\text{CRB}_{\text{NU}}^{\text{CG}}(\theta_1)} \) as a function of the noncircularity rate \( \rho_1 \) for different values of the SNR.

Fig. 4. Ratio \( r \triangleq \frac{\text{CRB}_{\text{NU}}^{\text{NGC}}(\theta_1)}{\text{CRB}_{\text{NU}}^{\text{CG}}(\theta_1)} \) as a function of the DOA separation for different values of DOA separation with SNR = -5 dB and \( \Delta \phi = 0.52 \) rd.

It shows that \( \text{CRB}_{\text{NU}}^{\text{NGC}}(\theta_1) \) decreases monotonically as the noncircularity rate increases but is relatively insensitive to this increase except for very low SNR (\( r_1 \leq ||a||^2 \)).

Figs. 4 and 5 consider two sources and examine the dependence of the ratio \( r \) with the noncircularity rate \( \rho_2 = \rho_1 \), the DOA separation \( \Delta \theta = \theta_2 - \theta_1 \) and the SNR. Fig. 4 shows that \( \text{CRB}_{\text{NU}}^{\text{NGC}}(\theta_1) \) decreases as the noncircularity rate increases (this extends to two equipowered and independent sources a property proved in the one source case). Furthermore this decrease is more prominent for low DOA separation. Fig. 5 shows that \( r \) decreases as the DOA separation and the SNR decreases and the difference of order of magnitude between \( \text{CRB}_{\text{NU}}^{\text{NGC}}(\theta_1) \) and \( \text{CRB}_{\text{NU}}^{\text{CG}}(\theta_1) \) is quite significant for low DOA separations and SNRs.

VI. CONCLUSION

In this paper, the stochastic CRBs of the DOA parameter estimates for noncircular complex Gaussian sources in the general case of an arbitrary unknown Gaussian noise field is considered. New closed-form expressions of these bounds have been presented. Compared with the deterministic CRB and the circular complex Gaussian CRB, some properties have been proved and some numerical examples with particular noise fields have been exhibited. They show that the difference between the noncircular and circular complex Gaussian CRB may be quite significant, particularly for low DOA separations and SNR’s. Consequently our derived noncircular complex Gaussian CRB provides a tighter upper bound on the CRB under noncircular complex discrete distribution compared to the standard circular complex Gaussian CRB.

APPENDIX I

PROOF OF LEMMA 1

Maximizing the log-likelihood (3.1) without any constraint on the Hermitian matrix \( \mathbf{R}_S \) reduces to a standard maximization problem, whose derivation follows the steps in [12]. A necessary condition for an extremum of (3.1) with respect to \( \mathbf{R}_S \) is...
that the partial derivatives of (3.1) with respect to the entries of $R_s$ be equated to zero. Jaffee [12] shows that this condition is equivalent to [12, rel. (4)], which is preserved here, where $R_y$ is replaced by $R_\theta$. It yields here

$$
\hat{A}^H(\theta)R_\theta^{-1}(R_y, T - R_y)R_y^{-1}\hat{A}(\theta) = 0. \quad (A.1)
$$

In order to obtain an expression of $\hat{R}_{s,ML}$, we need an expression for the inverse of $R_\theta$. We omit the dependence of $A(\theta)$ on $\theta$ to simplify the notation and define

$$
\hat{A} \triangleq Q_n^{-1/2}A \quad \text{with} \quad \hat{A} \triangleq Q_n^{-1/2}\hat{A} = \begin{bmatrix} \hat{A} & O \\ O & \hat{A}^* \end{bmatrix}
$$

where $\hat{A} \triangleq Q_n^{-1/2}A$. From the matrix inversion lemma (see e.g., [2, Lemma A.3]), we obtain

$$
\hat{R}_\theta^{-1} = I_{2M} - \hat{A}(I_{2K} + R_s\hat{A}^H\hat{A})^{-1}R_s\hat{A}^H \quad \text{and} \quad
$$

$$
R_\theta^{-1} = Q_n^{-1} - Q_n^{-1/2}\hat{A}(I_{2K} + R_s\hat{A}^H\hat{A})^{-1}R_s\hat{A}^HQ_n^{-1/2}
$$

and after some straightforward calculations

$$
R_{\theta}^{-1}\hat{A} = Q_n^{-1/2}\hat{A}(I_{2K} + R_s\hat{A}^H\hat{A})^{-1}. \quad \text{(A.2)}
$$

Using this relation, the condition (A.1) becomes

$$
\hat{A}^HQ_n^{-1/2}(R_y, T - R_y)Q_n^{-1/2}\hat{A} = 0. \quad \text{(A.2)}
$$

Inserting the expression of $R_\theta$ in (A.2) we get

$$
\hat{A}^HQ_n^{-1/2}R_y, TQ_n^{-1/2}\hat{A} = \hat{A}^HAR_s\hat{A}^H\hat{A} + \hat{A}^H\hat{A}. \quad \text{(A.3)}
$$

Consequently, the ML estimates of $R_s$ is given by

$$
\hat{R}_{s,ML} = [\hat{A}^H\hat{A}]^{-1}[\hat{A}^H\hat{R}_{y, T}\hat{A} - \hat{A}^H\hat{A}][\hat{A}^H\hat{A}]^{-1}
$$

with

$$
\hat{R}_{y, T} \triangleq Q_n^{-1/2}R_y, TQ_n^{-1/2}
$$

Because $\hat{R}_{y, T}$ and $[\hat{A}^H\hat{A}]^{-1}$ are partitioned as

$$
[\hat{A}^H\hat{A}]^{-1} = \begin{bmatrix} \hat{A}^H\hat{A}^{-1} & 0 \\ O & \hat{A}^T\hat{A}^* \end{bmatrix}
$$

with $\hat{R}_{y, T} \triangleq Q_n^{-1/2}R_y, TQ_n^{-1/2}$ and $\hat{R}_{y, T}' \triangleq Q_n^{-1/2}R_y, TQ_n^{-1/2}$, the unconstrained ML estimates of $R_s$ (A.3) is partitioned as well. Consequently, the joint constrained and unconstrained ML estimate of $(\theta, \rho, \sigma)$ coincide.

**APPENDIX II
PROOF OF RESULT 1**

Based on the Fisher information matrix given here by the extended Slepian–Bangs formula derived in [10]

$$(I_F)_{k, l} = \frac{1}{2} \text{Tr} \left[ \frac{\partial R_\theta}{\partial \sigma_k} R_\theta^{-1} \frac{\partial R_\theta}{\partial \sigma_l} R_\theta^{-1} \right]$$

all the steps of the derivation of [7] apply where the [7, rel. (19)] is replaced by

$$
R_s = [c_1, \ldots, c_K] = \begin{bmatrix} c_1^H \\ \vdots \\ c_K^H \\ c_1' \\ \vdots \\ c_K' \end{bmatrix}
$$

and

$$
R_s' = [c_1', \ldots, c_K'] = \begin{bmatrix} c_1 \\ \vdots \\ c_K \end{bmatrix}
$$

Then, [7, rels. (20) and (22)] become, respectively

$$
\frac{\partial R_\hat{\theta}}{\partial \sigma_k} = D_k C_k^H \hat{A}^H + \hat{A}C_k D_k^H \quad \text{with} \quad D_k \triangleq \begin{bmatrix} d_k & 0 \\ 0 & d_k' \end{bmatrix}
$$

$$(F)_{k, l} = 2R \left\{ \text{Tr} \left[ (D_k^H \Pi_{\hat{\theta}} D_l) (C_l^H \hat{A}^H R_{\theta}^{-1} A C_k) \right] \right\} \quad \text{(B.4)}
$$

with

$$
(F)_{k, l} = 2R \left\{ \text{Tr} \left[ (D_k^H \Pi_{\hat{\theta}} A Q_n^{-1/2} A C_k) \right] \right\}
$$

$$
(M_{\hat{\theta}})_{k, l} = 2R \left\{ \text{Tr} \left[ (D_k^H \Pi_{\hat{\theta}} Q_n^{-1/2} C_k) \right] \right\}
$$

$$
(T_{\hat{\theta}})_{k, l} = 2R \left\{ \text{Tr} \left[ Q_n^{-1/2} A C_k^H C_l^H \right] \right\}
$$

$$
 (\Pi_{\hat{\theta}})_{k, l} = 2R \left\{ \text{Tr} \left[ Q_n^{-1/2} A C_k^H C_l^H \right] \right\}
$$

and

$$
 Q_n^{-1/2} C_k^H C_l^H = \begin{bmatrix} Q_{k, l} & O \\ O & Q_{k, l} \end{bmatrix}
$$

$$
 \frac{\partial Q_n}{\partial \sigma_k} = Q_n^{-1/2} C_k^H C_l^H = \begin{bmatrix} Q_{k, l} & O \\ O & Q_{k, l} \end{bmatrix}
$$

Then, $Q_n^{-1/2}C_k^H C_l^H = Q_{k, l}$ and $\hat{A} \triangleq Q_n^{-1/2}A = \begin{bmatrix} \hat{A} & O \\ O & \hat{A}^* \end{bmatrix}$ where

$$
Q_n^{-1/2}C_k^H C_l^H = \begin{bmatrix} G_k & G_k^* \\ G_k^* & G_k \end{bmatrix}
$$

and $\Pi_{\hat{\theta}}$, we obtain the first equation shown at the bottom of the next page. This gives, after some straightforward calculations

$$
\textbf{F} = 4R \left[ (D_k^H \Pi_{\hat{\theta}} D_l) \circ (R_s \hat{A}^H R_s' \hat{A}^T \hat{R}_{y, T}' \hat{R}_{y, T}^{-1} \hat{A} R_s) \right]. \quad \text{(B.7)}
$$
Finally, adapting the same calculations used to derive the closed-form expression of $F$, to matrices $M_{NU}^{NG}$ (B.5) and $T_{AU}^{NG}$ (B.6), [7, rels. (43) and (44)] are replaced by (3.3) and (3.4), respectively.

**APPENDIX III**

**Proof of Result 2**

All the steps of the derivation of [7, Section V.1] apply where here

$$Q_n = Q_{n-1}^{-1/2} \left( \begin{array}{ccc} e_k & e_k^T & \mathbf{0} \\ \mathbf{0} & \mathbf{0} & e_k^T \end{array} \right) Q_{n-1}^{1/2}$$

$$C_k = R_{\theta} \left( \begin{array}{c} e_k \\ 0 \\ e_k \end{array} \right)$$

with $D = \left[ d_1, \ldots, d_K \right]$. Inserting this into (B.5) and (B.6), [7, rels. (54) and (56)] and [7, rel. (55)] are replaced, respectively, by the second equation shown at the bottom of the page. Following the last steps of [7, Sec. V.1] and rewriting [7, rel. (59)] as $M_{NU}^{NG} = MQ_{n-1}^{1/2}$ and $T_{NU}^{NG} = Q_{n-1}^{1/2}TQ_{n-1}^{-1}$, Result 2 is proved.

**APPENDIX IV**

**Proof of Result 3**

In the single source case, the two covariance matrices must be rewritten as

$$R_y = \sigma_n^2 \mathbf{a}_1^H \mathbf{a}_1 + Q_n \quad \text{and} \quad R'_y = \sigma_n^2 \rho \mathbf{a}_1^H \mathbf{a}_1^T \mathbf{a}_1$$

Consequently thanks to the matrix inversion lemma, we get

$$G = \mathbf{I}_M - \frac{\beta}{1 + \beta \mathbf{a}_1^H \mathbf{a}_1} \mathbf{a}_1 \mathbf{a}_1^T$$

and

$$G' = \frac{\sigma_n^2 \rho \mathbf{a}_1^H \mathbf{a}_1}{(1 + \sigma_n^2 (1 + \rho^2) \mathbf{a}_1^H \mathbf{a}_1)} \mathbf{a}_1 \mathbf{a}_1^T$$

(D.8)

where $\mathbf{a}_1 = Q_{n-1}^{-1/2} \mathbf{a}_1$ and where $\beta \triangleq \sigma_n^2 (1 + \sigma_n^2 (1 + \rho^2) \mathbf{a}_1^H \mathbf{a}_1) / (1 + \sigma_n^2 \mathbf{a}_1^H \mathbf{a}_1)$ is a real constant. In the case of nonuniform white noise field, the proof follows along the lines of [6, App. D]. Using $d_k = \mathbf{a}_1 \mathbf{a}_1^T$, with $\mathbf{A}$ is a real-valued diagonal matrix, the matrix $M_{NU}^{NG}$ (3.6) becomes

$$M_{NU}^{NG} = 2R \left[ (\sigma_n^2 \mathbf{G} \mathbf{a}_1 + \mathbf{A}^H \mathbf{a}_1^T) \odot (\mathbf{d}_k^H \mathbf{I}_{n-1}) \right]$$

$$= 2R \left[ \frac{i \sigma_n^2}{1 + \sigma_n^2 (1 + \rho^2) \mathbf{a}_1^H \mathbf{a}_1} \mathbf{a}_1^T \odot \left( \mathbf{A}^H (\mathbf{A} - \mathbf{A}^H) \mathbf{a}_1 \mathbf{a}_1^T \mathbf{I}_M \right) \right]$$

because $\mathbf{A} = (\mathbf{A}^H) \mathbf{I}_{n-1}^2$ is a real-valued diagonal matrix and consequently $\mathbf{A}^T \odot (\mathbf{A}^H (\mathbf{A} - \mathbf{A}^H) \mathbf{a}_1 \mathbf{a}_1^T) \mathbf{I}_{n-1}$ is a real row vector. Using the structure of $R_{\theta}^2$ and (D.8), (3.8) follows thanks to straightforward but tedious calculations.

**APPENDIX V**

**Proof of Lemma 2**

Based on the assumptions of the lemma, the Jacobian matrix $D_{\theta}^{A_k}$ of the mapping that associates the estimate $\theta_T$ to $R_{y,T}$ satisfies the constraint (see [15])

$$D_{\theta}^{A_k} (A^* \otimes A) = \mathbf{O}$$

and because the covariance matrix $\text{Cov}_{r_y}$ of the asymptotic distribution of $\text{vec}(R_{y,T})$ is given by simple algebraic manipulations of $\text{Cov}_{r_y} = \mathbb{E} \left( \text{vec} (R_{y,T}) - \text{vec} (R_{y,A}) \right)^2$ which is a simple extension of the expression of $\text{Cov}_{r_y}$ given in [16] for $q_n = \sigma_n^2 \mathbf{I}_M$ and for an arbitrary second-order distribution:

$$\text{Cov}_{r_y} = (A^* \otimes A) \text{Cov}_{r_y} (A^T \otimes A^H) + Q^* \otimes Q_n$$

$$+ Q^* \otimes \mathbf{AR}^* \mathbf{A}^H + A^* \mathbf{R}^* \mathbf{A}^T \otimes Q_n$$

with $\text{Cov}_{r_y} = R_{\theta}^* \otimes R_{\theta} + K(R_r^* \otimes R_{\theta}^*)$, the first term of $\text{Cov}_{r_y}$ which contains $R_{\theta}^*$ disappears in the expression of the covariance $\text{Cov}_{\theta}^{A_k} = D_{\theta}^{A_k} \text{Cov}_{r_y} \left( D_{\theta}^{A_k} \right)^T$ of the asymptotic distribution of $\theta_T$. 

\[
(F)_{k,t} = 2R \left[ \text{Tr} \left( \left( \begin{array}{ccc} d_k^H & 0 \\ 0 & d_k^T \end{array} \right) \left( \begin{array}{ccc} \Pi_{\mathbf{A}}^{1/2} & 0 \\ 0 & \Pi_{\mathbf{A}}^{-1/2} \end{array} \right) \left( \begin{array}{c} d_t \\ 0 \end{array} \right) \right) \left( \begin{array}{ccc} e_k^H & e_k^T \\ e_k^T & e_k \end{array} \right) \left( \begin{array}{ccc} \Pi_{\mathbf{A}}^{1/2} & 0 \\ 0 & \Pi_{\mathbf{A}}^{-1/2} \end{array} \right) \left( \begin{array}{c} d_t \\ 0 \end{array} \right) \right) \right]
\]

\[
(M)_{k,t} = 2R \left[ \text{Tr} \left( \left( \begin{array}{ccc} e_k^H \Pi_{\mathbf{A}}^{1/2} & 0 \\ 0 & e_k \end{array} \right) \left( \begin{array}{ccc} \Pi_{\mathbf{A}}^{1/2} & 0 \\ 0 & \Pi_{\mathbf{A}}^{-1/2} \end{array} \right) \left( \begin{array}{c} Q_n^{-1/2} e_k \left( Q_n^{-1/2} e_k^T Q_n^{-1/2} \right) \\ 0 \end{array} \right) \right) \left( \begin{array}{ccc} e_k^H & e_k^T \\ e_k^T & e_k \end{array} \right) \left( \begin{array}{ccc} \Pi_{\mathbf{A}}^{1/2} & 0 \\ 0 & \Pi_{\mathbf{A}}^{-1/2} \end{array} \right) \left( \begin{array}{c} Q_n^{-1/2} e_k \left( Q_n^{-1/2} e_k^T Q_n^{-1/2} \right) \\ 0 \end{array} \right) \right) \right]
\]

\[
(T)_{k,t} = 4R \left[ \text{Tr} \left( \left( \begin{array}{ccc} e_k^H & 0 \\ 0 & e_k \end{array} \right) \left( \begin{array}{ccc} \Pi_{\mathbf{A}}^{1/2} & 0 \\ 0 & \Pi_{\mathbf{A}}^{-1/2} \end{array} \right) \left( \begin{array}{c} Q_n^{-1/2} e_k \left( Q_n^{-1/2} e_k^T Q_n^{-1/2} \right) \\ 0 \end{array} \right) \right) \left( \begin{array}{ccc} e_k^H & e_k^T \\ e_k^T & e_k \end{array} \right) \left( \begin{array}{ccc} \Pi_{\mathbf{A}}^{1/2} & 0 \\ 0 & \Pi_{\mathbf{A}}^{-1/2} \end{array} \right) \left( \begin{array}{c} Q_n^{-1/2} e_k \left( Q_n^{-1/2} e_k^T Q_n^{-1/2} \right) \\ 0 \end{array} \right) \right) \right]
\]
APPENDIX VI

PROOF OF RESULT 5

Applying [2, rel. (4.3)] to the relation $\mathbf{r}_g = \mathbf{A} \mathbf{s} \mathbf{A}^H + \mathbf{l}_{2M}$, we have

$$\mathbf{R}_s - \mathbf{R}_s \mathbf{A}^H \mathbf{r}_g^{-1} \mathbf{A}^* \mathbf{R}_s = (\mathbf{R}_s^{-1} + \mathbf{A}^H \mathbf{A})^{-1}$$

and the left upper $K \times K$ block of (F.9)

$\mathbf{R}_s - \mathbf{R}_s \mathbf{A}^H \mathbf{r}_g^{-1} \mathbf{A}^* \mathbf{R}_s$ is positive definite as well.

Consequently $\text{CRB}_{\text{DET}}(\theta) \leq 2\mathbf{F}^{-1}$ thanks to (4.5) and (B.7). Because $\mathbf{T}^{-1}$ is positive definite (see [7, rel. (18)]), Result 4 follows from $\mathbf{F}^{-1} \leq (\mathbf{F} - \mathbf{M} \mathbf{T}^{-1} \mathbf{M}^T)^{-1}$ applied to (B.4).

In the particular case where the sources are uncorrelated with noncircularity rate equals to 1, $\mathbf{R}_s' = \mathbf{R}_s \mathbf{A}^*$ where $\mathbf{R}_s$ and $\mathbf{A}$ are diagonal with $\Phi = \text{Diag}(e^{j\phi_1}, \ldots, e^{j\phi_N})$ and

$$[\mathbf{R}_s \mathbf{A}^H, \mathbf{R}_s \mathbf{A}^T] \mathbf{r}_g^{-1} = \mathbf{R}_s \left[ \mathbf{A}^H, \mathbf{A}^* \right] \mathbf{r}_g^{-1} \left[ \mathbf{A}^H, \mathbf{A}^* \right] \mathbf{R}_s.$$

Result 4 extends by applying now [2, rel. (4.3)] to the relation

$$\mathbf{r}_g = \left[ \mathbf{A}^H, \mathbf{A}^* \right] \mathbf{R}_s \mathbf{A}^H + \mathbf{l}_{2M}.$$

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